

**Homework 2: Solutions**

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1. (a)

Global Balance Equations:

$$\begin{aligned} \pi_1 p_1 + \pi_1(1 - p_1) &= \pi_2(1 - p_2) + \pi_3 p_3 \\ \pi_2 p_2 + \pi_2(1 - p_2) &= \pi_3(1 - p_3) + \pi_1 p_1 \\ \pi_3 p_3 + \pi_3(1 - p_3) &= \pi_1(1 - p_1) + \pi_2 p_2 \\ \pi_1 + \pi_2 + \pi_3 &= 1 \end{aligned}$$

Then:

$$\begin{aligned} \pi_1 &= \frac{1 - p_2(1 - p_3)}{(2 - p_1)(1 - p_2(1 - p_3)) + (1 + p_2)(1 - p_3(1 - p_1))} \\ &= \frac{1 - p_2 + p_2 p_3}{3 - p_1 - p_2 - p_3 + p_1 p_2 + p_1 p_3 + p_2 p_3} \end{aligned}$$

$\pi_2$  and  $\pi_3$  can be written similarly.

(b)

$\{A\} = \{\text{a clockwise move that is then followed by one consecutive counterclockwise move}\}$

$$\mathbb{P}(A) = \sum_{i=1}^3 \mathbb{P}(A|S_i) \pi_i = p_1(1 - p_2)\pi_1 + p_2(1 - p_3)\pi_2 + p_3(1 - p_1)\pi_3$$

2. (a)

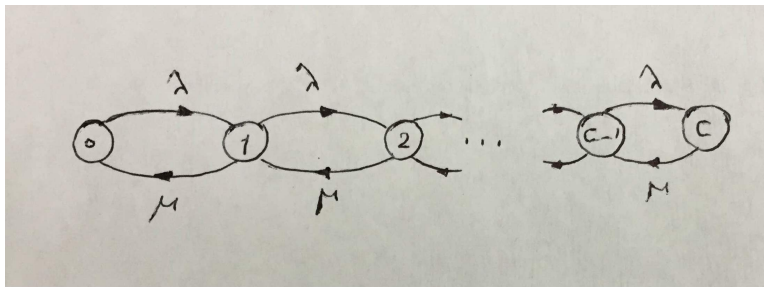


Figure 1: The Markov Chain of M/M/1/C Queue

Detailed Balance Equations:

$$\begin{aligned} \pi_0 \lambda &= \pi_1 \mu \\ \pi_1 \lambda &= \pi_2 \mu \\ &\dots \\ \pi_{c-1} \lambda &= \pi_c \mu \\ \sum_{i=0}^c \pi_i &= 1 \end{aligned}$$

So

$$\pi_i = \left(\frac{\lambda}{\mu}\right)^i \pi_0$$
$$\sum_{i=0}^c \left(\frac{\lambda}{\mu}\right)^i \pi_0 = 1$$

Case 1:  $\lambda = \mu$

$$\sum_{i=0}^c \left(\frac{\lambda}{\mu}\right)^i \pi_0 = 1 \Rightarrow \pi_0 = \frac{1}{c+1}$$
$$\pi_i = \pi_0, \text{ for } i=0 \text{ to } c$$

Case 2:  $\lambda \neq \mu$

$$\rho = \frac{\lambda}{\mu}:$$

$$\sum_{i=0}^c \left(\frac{\lambda}{\mu}\right)^i \pi_0 = 1 \Rightarrow \pi_0 = \frac{1-\rho}{1-\rho^{c+1}}$$
$$\pi_i = \rho^i \frac{1-\rho}{1-\rho^{c+1}}, \text{ for } i=0 \text{ to } c$$

This system is stable for any  $\rho$ .

(b)

N: Number of customers in the system.

$$\mathbb{E}[N] = \sum_{i=0}^c i \pi_i$$

Case 1:  $\lambda = \mu$

$$\mathbb{E}[N] = \frac{1}{c+1} \times \frac{c(c+1)}{2} = \frac{c}{2}$$

Case 2:  $\lambda \neq \mu$

$$\rho = \frac{\lambda}{\mu}:$$

$$\mathbb{E}[N] = \sum_{i=0}^c i \rho^i \frac{1-\rho}{1-\rho^{c+1}}$$
$$= \frac{\rho(1-\rho)}{1-\rho^{c+1}} \times \frac{-(c+1)\rho^c(1-\rho) + 1 - \rho^{c+1}}{(1-\rho)^2}$$
$$= \frac{\rho}{1-\rho} - \frac{(c+1)\rho^{c+1}}{1-\rho^{c+1}}$$

\* To find out  $\sum_{i=0}^c i \rho^i$ :

$$\sum_{i=0}^c \rho^i = \frac{1-\rho^{c+1}}{1-\rho}$$

Then we take the derivative of both sides with respect to  $\rho$ :

$$\sum_{i=0}^c i\rho^{1-i} = \frac{-(c+1)\rho^c(1-\rho) + 1 - \rho^{c+1}}{(1-\rho)^2}$$

So  $\sum_{i=0}^c i\rho^i = \rho \times \sum_{i=0}^c i\rho^{1-i} = \rho \times \frac{-(c+1)\rho^c(1-\rho) + 1 - \rho^{c+1}}{(1-\rho)^2}$ .

(c)

When a customer sees the system in state  $c$ , he/she cannot enter the system. By PASTA, drop probability is equal to  $\pi_c$ . So the effective rate of arrival for this system is  $\lambda(1 - \pi_c)$ . Then by Little's law:

$$\lambda_{\text{eff}} \times \mathbb{E}[D] = \mathbb{E}[N] \longrightarrow \mathbb{E}[D] = \frac{\mathbb{E}[N]}{\lambda(1 - \pi_c)}$$

3. (a)

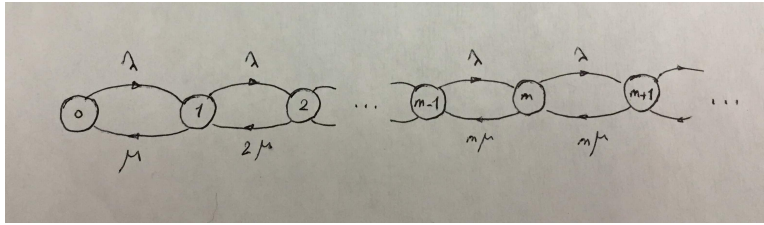


Figure 2: The Markov Chain of M/M/∞ Queue

Detailed Balance Equations:

$$\begin{aligned} \sum_{i=0}^{\infty} \pi_i &= 1 \\ \lambda\pi_0 &= \mu\pi_1 \\ \lambda\pi_1 &= 2\mu\pi_2 \\ &\dots \\ \lambda\pi_{m-1} &= m\mu\pi_m \\ \lambda\pi_m &= m\mu\pi_{m+1} \\ &\dots \end{aligned}$$

Let  $\rho = \frac{\lambda}{\mu}$ , then:

$$\begin{aligned} \pi_i &= \frac{\rho^i}{i!} \pi_0, \text{ for } i=0 \text{ to } m \\ \pi_i &= \left(\frac{\rho}{m}\right)^{i-m} \pi_m = \left(\frac{\rho}{m}\right)^{i-m} \frac{\rho^m}{m!} \pi_0, \text{ for } i=m+1, m+2, \dots \end{aligned}$$

Then to find  $\pi_0$ :

$$\begin{aligned} \sum_{i=0}^{\infty} \pi_i &= 1 \Rightarrow \sum_{i=0}^m \frac{\rho^i}{i!} \pi_0 + \sum_{i=m+1}^{\infty} \left(\frac{\rho}{m}\right)^{i-m} \frac{\rho^m}{m!} \pi_0 = 1 \\ \Rightarrow \pi_0 &= \frac{1}{\sum_{i=0}^m \frac{\rho^i}{i!} + \frac{\rho^{m+1}}{m \times m!} \frac{1}{1 - \frac{\rho}{m}}} \end{aligned}$$

\* For  $\rho < m$ , which is the stability condition.

(b)

$$\begin{aligned}\mathbb{P}(\text{a customer has to wait}) &= \mathbb{P}(\text{a customer sees } m \text{ customer or more in the system}) \\ &\stackrel{*}{=} \mathbb{P}(m \text{ persons or more in the system}) \\ &= \sum_{i=m}^{\infty} \pi_i = 1 - \sum_{i=0}^{m-1} \pi_i = 1 - \pi_0 \sum_{i=0}^{m-1} \frac{\rho^i}{i!}\end{aligned}$$

\* By using PASTA.

(c)

$$\pi_0 = \frac{1}{\sum_{i=0}^{\infty} \frac{\rho^i}{i!}} = e^{-\rho} \longrightarrow \pi_i = e^{-\rho} \frac{\rho^i}{i!}$$

Note that this is a Poisson distribution with mean  $\rho$ . Hence, the number of customers in the  $M/M/\infty$  system in steady state is simply a Poisson random variable with mean  $\rho = \frac{\lambda}{\mu}$ .

Also, the system is stable for all  $\rho < \infty$ . And,  $\mathbb{P}(\text{a customer has to wait}) \longrightarrow 0$

4.  $N(t) \sim \text{Poisson}(\alpha)$ ,  $\mathbb{P}(\text{transmission}) = \frac{1}{\alpha}$

(a)

$$\begin{aligned}\mathbb{P}(\text{slot is idle}) &= \mathbb{P}(\text{no packet is transmitted}) \\ &= \sum_{n=0}^{\infty} \mathbb{P}(\text{no packet is transmitted} | N(t) = n) \mathbb{P}(N(t) = n) \\ &= \sum_{n=0}^{\infty} \left(1 - \frac{1}{\alpha}\right)^n e^{-\alpha} \frac{\alpha^n}{n!} = e^{-\alpha} \sum_{n=0}^{\infty} \frac{(\alpha - 1)^n}{n!} \\ &= e^{-\alpha} e^{\alpha-1} = \frac{1}{e}\end{aligned}$$

(b)

$$\begin{aligned}\mathbb{P}(N = n | \text{slot is idle}) &= \frac{\mathbb{P}(\text{slot is idle} | N = n) \mathbb{P}(N = n)}{\mathbb{P}(\text{slot is idle})} \\ &= \frac{\left(1 - \frac{1}{\alpha}\right)^n e^{-\alpha} \frac{\alpha^n}{n!}}{\frac{1}{e}} = e^{-(\alpha-1)} \frac{(\alpha-1)^n}{n!} \\ &\sim \text{Poisson}(\alpha - 1)\end{aligned}$$

(c)

$$\begin{aligned}\mathbb{P}(\text{slot is successful}) &= \mathbb{P}(\text{only one station transmits}) \\ &= \sum_{n=1}^{\infty} \mathbb{P}(\text{only one station transmits} | N(t) = n) \mathbb{P}(N(t) = n) \\ &= \sum_{n=1}^{\infty} \binom{n}{1} \frac{1}{\alpha} \left(1 - \frac{1}{\alpha}\right)^{n-1} e^{-\alpha} \frac{\alpha^n}{n!} = \sum_{n=1}^{\infty} e^{-\alpha} \frac{(\alpha-1)^{n-1}}{(n-1)!} \\ &= e^{-1} \sum_{n=0}^{\infty} e^{-(\alpha-1)} \frac{(\alpha-1)^n}{n!} = \frac{1}{e}\end{aligned}$$

(d)

$$\begin{aligned}\mathbb{P}(N = n + 1 | \text{slot is successful}) &= \frac{\mathbb{P}(\text{slot is successful} | N = n + 1) \mathbb{P}(N = n + 1)}{\mathbb{P}(\text{slot is successful})} \\ &= \frac{(n + 1) \frac{1}{\alpha} \left(1 - \frac{1}{\alpha}\right)^n e^{-\alpha} \frac{\alpha^{n+1}}{(n+1)!}}{\frac{1}{e}} = e^{-(\alpha-1)} \frac{(\alpha - 1)^n}{n!} \\ &\sim \text{Poisson}(\alpha - 1)\end{aligned}$$